

The Binomial Case

It can be shown that the influence statistic has a non-central chi-squared distribution with 2 d.f. with non-centrality parameter λ given by

$$\lambda = 0.5 * \left[\frac{(E(\hat{p}_1 - \hat{p}_{1(-i)}))^2}{V_{11}} + \frac{(E(\hat{p}_2 - \hat{p}_{2(-i)}))^2}{V_{22}} \right],$$

for $E(\hat{p}_1 - \hat{p}_{1(-i)}) = \frac{N_{1i}(p_{1i} - p_{1(-i)})}{N_{1(-i)} + N_{1i}}$, a similar expression for $E(\hat{p}_2 - \hat{p}_{2(-i)})$,

$$V_{11} = Var(\hat{p}_1 - \hat{p}_{1(-i)}) = \frac{N_{1i}p_{1i}(1-p_{1i})}{(N_{1(-i)} + N_{1i})^2} + \frac{N_{1i}^2 p_{1(-i)}(1-p_{1(-i)})}{(N_{1(-i)} + N_{1i})^2 N_{1(-i)}}$$
 and a similar expressions for V_{22} ,

where, \hat{p}_1 is the proportion of responders in the first arm and $\hat{p}_{1(-i)}$ is the proportion of responders in the first arm after excluding the i^{th} possibly influential stratum. $N_{1(-i)}$ and N_{1i} are the number of subjects in the first arm for those outside stratum i and those in stratum i respectively. \hat{p}_2 , $\hat{p}_{2(-i)}$, $N_{2(-i)}$ and N_{2i} are similarly defined for the second arm.

The scaled inflation in influence is then given by λ , the percent inflation in influence by $100*\lambda$ and the percent inflation in variance of the influence statistic is given by 200λ . These follow from the expected value of a central chi-square with n d.f. given as n and variance given as 2n. An aberrant large site/strata will lead to an influence statistic with a non-central distribution with expected value $n+2\lambda$ and variance $2(n+4\lambda)$.

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The Continuous Case

It can be shown that the influence statistic has a non-central chi-squared distribution with 2 d.f. with non-centrality parameter λ given by

$$\lambda = 0.5 * \left[\frac{(E(\bar{Y}_1 - \bar{Y}_{1(-i)}))^2}{V_{11}} + \frac{(E(\bar{Y}_2 - \bar{Y}_{2(-i)}))^2}{V_{22}} \right],$$

for $E(\bar{Y}_1 - \bar{Y}_{1(-i)}) = \frac{N_{1i}(\mu_{1i} - \mu_{1(-i)})}{N_{1(-i)} + N_{1i}}$, a similar expression for $E(\bar{Y}_2 - \bar{Y}_{2(-i)})$,

$$V_{11} = Var(\bar{Y}_1 - \bar{Y}_{1(-i)}) = \frac{N_{1i}\sigma_{1i}^2}{(N_{1(-i)} + N_{1i})^2} + \frac{N_{1i}^2\sigma_{1(-i)}^2}{(N_{1(-i)} + N_{1i})^2 N_{1(-i)}}, \text{ and a similar expression for } V_{22},$$

where, \bar{Y}_1 is the mean response in the first arm and $\bar{Y}_{1(-i)}$ is the mean response in the first arm after excluding the i^{th} possibly influential stratum. $N_{1(-i)}$ and N_{1i} are the number of subjects in the first arm for those outside stratum i and those in stratum i respectively. The mean and variance of the $N_{1(-i)}$ observations are denoted by $\mu_{1(-i)}$ and $\sigma_{1(-i)}^2$ and that for the N_{1i} observations are denoted by μ_{1i} and σ_{1i}^2 . \bar{Y}_2 , $\bar{Y}_{2(-i)}$, $N_{2(-i)}$ and N_{2i} , $\mu_{2(-i)}$, $\sigma_{2(-i)}^2$, μ_{2i} , and σ_{2i}^2 are defined similarly for the second arm.

The scaled inflation in influence is then given by λ , the percent inflation in influence by $100*\lambda$ and the percent inflation in variance of the influence statistic is given by 200λ . These follow from the expected value of a central chi-square with n d.f. given as n and variance given as 2n. An aberrant large site/strata will lead to an influence statistic with a non-central distribution with expected value $n+2\lambda$ and variance $2(n+4\lambda)$.

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The Survival Case

It can be shown that the influence statistic (the homogeneity of strata statistic- see Marubini E and Valsecchi MG. *Analyzing Survival Data from Clinical Trials and Observational Studies*) has a non-central chi-squared distribution with 1 d.f. with non-centrality parameter λ given by

$$\lambda = 0.5 * \left(V_{Li} (\xi_i - \bar{\xi})^2 + V_{L(-i)} (\xi_{-i} - \bar{\xi})^2 \right) \text{ with } \bar{\xi} = V_{Li} \xi_i + V_{L(-i)} \xi_{-i} / (V_{Li} + V_{L(-i)}).$$

The log hazard ratio in the i th and i -deleted strata are denoted by ξ_i and ξ_{-i} . The variances V_{Li} and $V_{L(-i)}$ can be obtained approximately as a the product $r*(1-r)$ of the randomization fractions r and $(1-r)$ multiplied by the expected number of events in the the i th and i -deleted strata (Schoenfeld, 1981, Bometrics). The variances V_{Li} is given as in the following (using expressions for the expected number of events from the Cytel EAST Manual.

$$V_{Li} = r(1-r) \left\{ \frac{a_{1i} \lambda_{1i}}{\lambda_{1i} + \gamma} \left[A - \frac{e^{-(\lambda_{1i} + \gamma)l}}{\lambda_{1i} + \gamma} \left(e^{(\lambda_{1i} + \gamma)A} - 1 \right) \right] + \frac{a_{2i} \lambda_{2i}}{\lambda_{2i} + \gamma} \left[A - \frac{e^{-(\lambda_{2i} + \gamma)l}}{\lambda_{2i} + \gamma} \left(e^{(\lambda_{2i} + \gamma)A} - 1 \right) \right] \right\}$$

for variable follow up. For the first and second arm in the i^{th} site, a_{1i} and a_{2i} are the uniform accrual rates and λ_{1i} and λ_{2i} are the hazard rates. A is the accrual period and γ is the drop-out hazard rate and l is the calendar time (A +follow-up). For fixed follow-up the expression for V_{Li} can be obtained on taking the limit as A approaches zero as

$$V_{Li} = r(1-r) \left\{ \frac{N_{1i} \lambda_{1i}}{\lambda_{1i} + \gamma} \left[1 - e^{-(\lambda_{1i} + \gamma)l} \right] + \frac{N_{2i} \lambda_{2i}}{\lambda_{2i} + \gamma} \left[1 - e^{-(\lambda_{2i} + \gamma)l} \right] \right\}$$

Similar expressions can be obtained for the $V_{L(-i)}$. The scaled inflation in influence is given by $\text{SQRT}(2)*\lambda$, the percent inflation in influence is 200λ and the percent inflation in variance is given by 400λ as the homogeneity of strata statistic is 1 d.f. chi-square.

Reference: Shankar Srinivasan and Arlene Swern. Measures of Expected Influence Provide Useful Constraints to Enrollment in Randomized Multi-Center Clinical Trials for Binomial, Continuous and Time-to-Event Endpoints, *Journal of Statistical Science and Application*, April 2015, Vol. 3, No. 3-4, 39-49.

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